

Statistics – Solutions

1. **D** – Of the choices, D is the only correct interpretation of a confidence interval.
2. **B** – The answer is the left endpoint of the confidence interval. Seen from the duality between confidence intervals and hypotheses testing.
3. **D** – Keep in mind that this is a one sided test. $Z = 2.79$, $p = 0.003$, reject at all reasonable levels of significance. Strong evidence that $\mu > 8$.
4. **E** – The answer is 1. Note that $f(x)$ never takes on negative values, so no knowledge of calculus is necessary. Formally: $P(X > 0) = 1 - P(X \leq 0) = 1 - \int_{-\infty}^0 f(x) dx = 1 - \int_{-\infty}^0 0 dx = 1$.
5. **C** – This can be seen after a simple integration; however, since calculus is not a prerequisite of statistics, we know that integration must not be required. We know that the area under a probability density function is always 1.
6. **D** – Because an easy test will have a high concentration of students scoring well, this distribution will be skewed to the left. In such a case, the mean will be below the median which will be below the mode. The only case in which this is true is choice D.
7. **A** – Definition of a Type II error.
8. **A** – A Chi-squared test for Goodness of Fit will be used, since we have an expectation and the observed counts are provided. The tricky part is determining the expected counts: There are possible combinations, listed as Player1, Player2. TT, TP, PT, TZ, ZT, PP, PZ, ZP, ZZ. Each of these should be equally likely to occur, with expected count $200/9 = 33.33$. However, notice that the matchups involving two different races are actually repeated (for example, TP and PT are both Terran/Protoss matchups). So the expected counts of matchups are as follows: TT – 33.33, TP – 66.67, TZ – 66.67, PP – 33.33, PZ – 66.67, ZZ – 33.33. Now, using the standard Chi-squared formula of $(O-E)^2/E$, the Chi-squared value becomes 3.29, which, using 5 degrees of freedom, provides a p-value greater than 0.3. The null hypothesis, that each matchup is represented in the correct proportions, cannot be rejected.
9. **D** – $P(X < (1377-1426)/\sigma) = .25$. $-49/\sigma = -.67$, so $\sigma = 73.134$ and $\sigma^2 = 5,348.6$.
10. **B** – $P(X < (1500-1426)/73.13) - P(X < (1400-1426)/73.13) = .8441 - .3611 = 0.4830$.
11. **C** – We first must create a new distribution for the herd of 200 cows: $\sim N(n\mu, \sqrt{n}\sigma) = N(285,200, 1034.2)$. Then, we know that $P(X < (x-285200)/1034.2) \leq .05$, so $Z \leq -1.645$, meaning that $x \leq 283498.7$. The largest whole number in that range is 283,498.
12. **C** – Either both D and E must work; or A, B, and C must all work. Since all components function independently, $P(D \text{ and } E) = p^2$, $P(A \text{ and } B \text{ and } C) = p^3$. $P(\text{either scenario}) = P(D,E) + P(A,B,C) = p^2 + p^3$. When $p = 0.6$, this expression evaluates to 0.58.
13. **D** – There are three total covariates. Just because two of them are highly correlated does not imply that you can ignore one of them. Therefore, you must use all three.
14. **D** – The transformation is to enlarge the response variables Y, so we could take e^Y .
15. **B** – Definition of Simpson's Paradox.
16. **A** – This is a question of conditional probability. Construct the following table for help with organization.

| | P | F | G | Total |
|----------------|-------------------|-------------------|-------------------|-------|
| D | $.01 * .3 = .003$ | $.01 * .4 = .004$ | $.01 * .3 = .003$ | .01 |
| D ^C | $.99 * .1 = .099$ | $.99 * .4 = .396$ | $.99 * .5 = .495$ | .99 |
| Total | .102 | .400 | .498 | 1 |

Then, the two answers become clear:

$$A - P(D/P) = P(P \& D) / P(P) = .003 / .102 = 0.0294$$

17. **D** – Using the same table as in question 16, $P(G) = 0.498$
18. **E** – By De Morgan’s identity, the complement of the intersection is the union of the complements, so the set is: $\{X < 3\} \cup \{X > 5\}$. There are an infinite number of natural numbers that lie in this set.
19. **C** – The expression for the total commute time is $X_1 + X_2 + X_3 + X_4 + X_5$. Each coefficient is 1, and the variance of each day’s time is 16. So the variance of the total weekly commute is $1^2*16 + 1^2*16 + 1^2*16 + 1^2*16 + 1^2*16 = 80$. Standard deviation = $\sqrt{80} = 8.94$
20. **D** – One square in each chart is shaded, representing $\frac{1}{4} * \frac{1}{4} = 1/16$, or $P(A)*P(B)=P(A \text{ and } B)$.
21. **B** – Correlation is not a resistant measure (it is highly influenced by outliers). Transformation of either variable changes the correlation. The higher the magnitude of a correlation, regardless of its sign, the stronger the relationship. However, it was not described as a linear relationship. Therefore, part C is false. A correlation of 0 means there is no *linear* relationship but there could easily be a strong non-linear relationship.
22. **D** – The hypotheses are $H_0: \mu \leq 1/2$, $H_A: \mu > 1/2$. Since the sample size is small, perform a t-test. The observed Z-value is $(.56-.5)/(0.05/(30^{.5})) = 6.573$. $P(t_{29} > 6.573) < 0.001$. Clearly, we should reject the null hypothesis at all reasonable significance levels.
23. **C** – The hypotheses are now $H_0: \mu_1 \leq \mu_2$, $H_A: \mu_1 > \mu_2$. The t-score is $t = \frac{X_1 - X_2 - (\mu_1 - \mu_2)}{\sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}} =$
- 1.80 for the given data. Using the conservative 29 degrees of freedom, the p-value lies between 0.01 and 0.05.
24. **B** – You could arrive 8:10-8:15 or 8:25-8:30 and have to wait under 5 minutes. This covers 10 of the 30 minutes, so the probability = $1/3$
25. **B** – The time you arrive is a uniformly distributed random variable X where x is between 0 and 30. If $0 < x < 15$, Y_1 is a uniform random variable taking values 0-15. If $15 < x < 30$, Y_2 is again a uniform variable taking values 0-15. So Y (taking realized values a) is a uniform distribution spanning from 0 to 15. The expected value of such a distribution is $E(Y) = (15-0)/2 = 7.5$
26. **D** – Written in order the numbers are: 1 1 2 3 4 5 5 6 9. The 5-number summary (Min, 1st quartile, Median, 3rd quartile, Max) is: 1, 1.5, 4, 5.5, 9. $9-1 = 8$, and $5.5-1.5 = 4$ so the ratio is 2.
27. **A** – The observed counts are given in the chart, and we already have expected counts based on Natalie’s hypothesis.
28. **A** – The null hypothesis for this test is $H_0: P_{\text{Cantor}} = P_{\text{Gauss}} = P_{\text{Poincaré}} = \frac{1}{4}$, $P_{\text{Galois}} = P_{\text{Gödel}} = P_{\text{Perelman}} = \frac{1}{12}$. The observed/expected counts are:

| | Cantor | Galois | Gauss | Gödel | Perelman | Poincaré |
|----------|--------|--------|-------|-------|----------|----------|
| Observed | 52 | 10 | 66 | 13 | 12 | 47 |
| Expected | 50 | 50/3 | 50 | 50/3 | 50/3 | 50 |

- There are 6-1, or 5 degrees of freedom in this test. The Chi-square statistic yields 10.16, corresponding to a p-value of 0.07, so we cannot reject the null at a 5% significance level.
29. **B** – Changing hypothesis after seeing the raw data is defined as “data snooping” or “data fishing”. It introduces bias and consequentially inflates the Type I error rate.
30. **B** – Let $x=N(0,2)$ and $y=N(0,2)$. $D^2/4 = (x^2+y^2)/4 = (x/2)^2 + (y/2)^2$. The distribution $x/2$ is represented by $N(0,1)$ and $y/2$ is represented by $N(0,1)$ as well. Thus we have $Z = X_1^2 + X_2^2$, where X_1 and X_2 are standard normal random variables, so $D^2/4$ is a Chi-square distribution with 2 degrees of freedom, which, by definition, has expected value 2.

